

[0,1] truncated lomax – lomax distribution with properties

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Abstract — We shall introduce a new continuous distribution dependent on $[0,1]$ truncated lomax ($[0,1]$ TLD) distribution , we will call it $[0,1]$ truncated lomax – lomax ($[0,1]$ TLLD) distribution. The some properties of the ($[0,1]$ TLLD) distribution will be derive , and also, the ($[0,1]$ TLLD) stress-strength model with different parameters will be obtain. The purpose of this work is to develop a new composite distribution that is more flexible than other distributions..

KEY WORDS — $[0,1]$ TLD, $[0,1]$ TLLD , Stress Strength Model of the $[0,1]$ TLLD.

I. INTRODUCTION

In recent times , Researchers attempted to suggest a new classes of statistical distributions. Here, we created a distribution we hope to benefit from it .This work will be on the generalizations stimulated by Eugene et al. [1] ,so that, they provided beta generalization distribution from (CDF) , $G(x)$ by.

$$F(x) = \left[\frac{1}{\beta(a,b)} \right] \int_0^{G(x)} Z^{a-1} [1-Z]^{b-1} dZ, \\ 0 < a, b < \infty \quad (1.1)$$

where $\beta(a,b) = \int_0^1 Z^{a-1} (1-Z)^{b-1} dZ$.

Jones [1][2] ,who used $X = G^{-1}(N)$ the arbitrary variable N for $N \sim \text{Beta}(a,b)$ and products X with cumulative distribution function (1.1) . Eugene et al , who knows a Beta normal (BN) distribution through placing $G(x)$ to be the cumulative distribution function of the normal distribution and also, inferred a few properties of this distribution while, Gupta and Nadarajah obtained formulas for the moments of above distribution [3]. Cordeiro and de Castro (2011) proposed Kumaraswamy generated family by exchanged the (BN) distribution with the (Kw) distribution.

The (PDF) relating to (1.1) is given by

$$f(x) = \frac{1}{\beta(a,b)} G(x)^{a-1} (1-G(x))^{b-1} g(x) \quad (1.2)$$

Where $g(x) = \frac{dG(x)}{dx}$ is (pdf) of the principle distribution.

The probability density and cumulative distribution functions of ($[0,1]$ TLD) are one by one

$$v(x) = \frac{\eta^\vartheta (1+\eta x)^{-(\vartheta+1)}}{(1-(1+\eta)^{-\vartheta})} \quad 0 < x < 1 \quad (1.3)$$

$$V(x) = \frac{(1-(1+\eta x)^{-\vartheta})}{(1-(1+\eta)^{-\vartheta})} \quad (1.4)$$

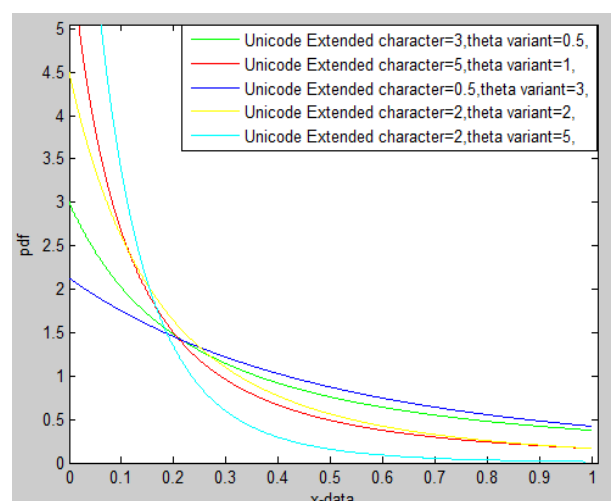


Figure 1: Plots PDF for ($[0,1]$ TLD) distribution .

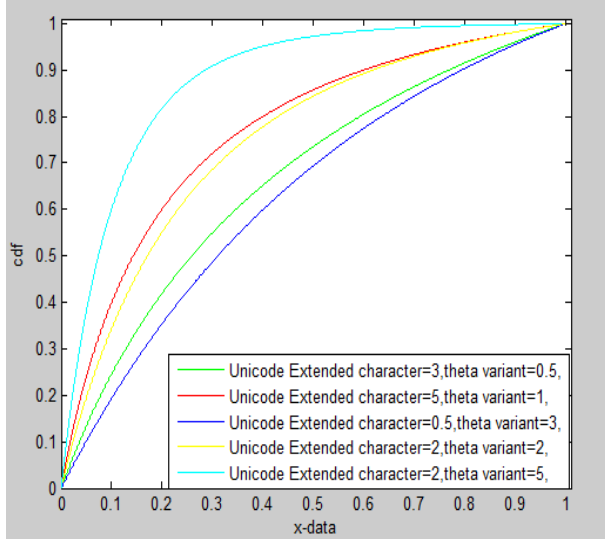


Figure 2: Plots CDF ([0,1] TLD) distribution

Presently, looking two totally cumulative distribution functions, W and G , our assume another distribution F by combination W with G , s.t $F(x) = W(G(x))$ be a cdf

$$F(x) = \int_0^{G(x)} \frac{\eta^\vartheta (1+\eta t)^{-(\vartheta+1)}}{(1-(1+\eta)^{-\vartheta})} dt$$

$$F(x) = \frac{1 - \{1 + \eta G(x)\}^{-\vartheta}}{1 - \{1 + \eta\}^{-\vartheta}} \quad (1.5)$$

and pdf :

$$f_s(x) = \frac{\partial}{\partial x} F(x) = \frac{\partial}{\partial x} \left[\frac{1 - \{1 + \eta G(x)\}^{-\vartheta}}{1 - \{1 + \eta\}^{-\vartheta}} \right]$$

$$f(x) = \frac{\eta^\vartheta \{1 + \eta G(x)\}^{-(\vartheta+1)}}{1 - \{1 + \eta\}^{-\vartheta}} \cdot g(x) \quad (1.6)$$

Where $g(x) = \frac{\partial G(x)}{\partial x}$

We will consider Eq (1.5) and (1.6) as a generalized class of distributions, its name ([0,1] TLD - G) distribution. A suppose G be Lomax distribution.

II. THE [0,1] TLLD DISTRIBUTION

A suppose that $g(x) = \kappa \lambda (1 + \kappa x)^{-(\lambda+1)}$ and $G(x) = 1 - (1 + \kappa x)^{-\lambda}$ referred as pdf and cdf for (r.v.) lomax, are one by one .at that point ,through the application (1.5) and (1.6) ,get the cumulative distribution and probability distribution functions for (r.v.) of ([0,1] TLLD) .

$$F(x) = \frac{1 - \{1 + \eta [1 - (1 + \kappa x)^{-\lambda}]\}^{-\vartheta}}{(1 - (1 + \eta)^{-\vartheta})} \quad (2.1)$$

$$f(x) = \frac{\eta^\vartheta \{1 + \eta [1 - (1 + \kappa x)^{-\lambda}]\}^{-(\vartheta+1)}}{(1 - (1 + \eta)^{-\vartheta})} \cdot \frac{\kappa \lambda (1 + \kappa x)^{-(\lambda+1)}}{(1 + \eta [1 - (1 + \kappa x)^{-\lambda}])^{-(\vartheta+1)}} \quad (2.2)$$

Furthermore, the reliability and the hazard rate functions are one by one

$$r(x) = 1 - F(x)$$

$$= 1 - \left[\frac{\{1 - (1 + \eta [1 - (1 + \kappa x)^{-\lambda}])^{-\vartheta}\}}{(1 - (1 + \eta)^{-\vartheta})} \right]$$

$$= \frac{(\{1 + \eta\}^{-\vartheta} + [1 + \eta (1 - \{1 + \kappa x\}^{-\lambda})]^{-\vartheta})}{(1 - (1 + \eta)^{-\vartheta})} \quad (2.3)$$

$$\gamma(x) = \frac{f(x)}{r(x)}$$

$$= \frac{\{\eta^\vartheta \kappa \lambda (1 + \kappa x)^{-(\lambda+1)} (1 + \eta [1 - (1 + \kappa x)^{-\lambda}])^{-(\vartheta+1)}\}}{\{\{1 + \eta\}^{-\vartheta} + (1 + \eta [1 - (1 + \kappa x)^{-\lambda}])^{-\vartheta}\}} \quad (2.4)$$

The r-th moment can be get by , [5] :

$$E(x^r) = \int_0^\infty x^r f(x) dx$$

$$E(x^r) = \int_0^\infty x^r \frac{\eta^\vartheta \kappa \lambda}{(1 - (1 + \eta)^{-\vartheta})} \cdot (1 + \kappa x)^{-(\lambda+1)} (1 + \eta [1 - (1 + \kappa x)^{-\lambda}])^{-(\vartheta+1)} dx$$

$$E(x^r) = \frac{\eta^\vartheta \kappa \lambda}{(1 - (1 + \eta)^{-\vartheta})} \int_0^\infty x^r (1 + \kappa x)^{-(\lambda+1)} (1 + \eta [1 - (1 + \kappa x)^{-\lambda}])^{-(\vartheta+1)} dx$$

Now , by using the series expansion

$$(1 - z)^{-n} = \sum_{\delta=0}^{\infty} \frac{\Gamma(n+\delta)}{\delta! \Gamma(n)} z^\delta \quad |z| < 1, \quad n > 0, \quad [6]$$

we get:

$$[1 + \eta (1 - \{1 + \kappa x\}^{-\lambda})]^{-(\vartheta+1)}$$

$$= [1 + \eta - \eta \{1 + \kappa x\}^{-\lambda}]^{-(\vartheta+1)}$$

$$[1 + \eta (1 - \{1 + \kappa x\}^{-\lambda})]^{-(\vartheta+1)} =$$

$$\sum_{\delta=0}^{\infty} \frac{\Gamma(\vartheta+1+\delta)}{\delta! \Gamma(\vartheta+1)} \{\eta [1 + \kappa x]^{-\lambda} - \eta\}^\delta$$

Thence

$$E(x^r) = \frac{\vartheta \eta \kappa \lambda}{[1 - \{1 + \eta\}^{-\vartheta}]} \int_0^\infty x^r \{1 + \kappa x\}^{-(\lambda+1)} \sum_{\delta=0}^\infty \frac{\Gamma(\vartheta+1+\delta)}{\delta! \Gamma(\vartheta+1)} [\eta \{1 + \kappa x\}^{-\lambda} - \eta]^\delta dx$$

$$\begin{aligned} E(x^r) &= \frac{\vartheta \eta \kappa \lambda}{[1 - \{1 + \eta\}^{-\vartheta}]} \sum_{\delta=0}^\infty \frac{\Gamma(\vartheta+1+\delta)}{\delta! \Gamma(\vartheta+1)} \int_0^\infty x^r \{1 + \kappa x\}^{-(\lambda+1)} [\eta \{1 + \kappa x\}^{-\lambda} - \eta]^\delta dx \\ &= \frac{\vartheta \eta \kappa \lambda}{[1 - \{1 + \eta\}^{-\vartheta}]} \sum_{\delta=0}^\infty \frac{\Gamma(\vartheta+1+\delta)}{\delta! \Gamma(\vartheta+1)} \int_0^\infty x^r \{1 + \kappa x\}^{-(\lambda+1)} [-\eta]^\delta (1 - (1 + \kappa x)^{-\lambda})^\delta dx \end{aligned}$$

Now , by applying

$$(1 - z)^b = \sum_{m=0}^\infty (-1)^m \frac{\Gamma(b+1)}{m! \Gamma(b-i+1)} z^m ,$$

$$|z| < 1, \quad b > 0, \quad [4] \quad (2.5)$$

we get:

$$[1 - \{1 + \kappa x\}^{-\lambda}]^\delta = \sum_{m=0}^\infty \frac{(-1)^m \Gamma(\delta+1)}{m! \Gamma(\delta-m+1)} [\{1 + \kappa x\}^{-\lambda}]^m$$

$$[1 - \{1 + \kappa x\}^{-\lambda}]^\delta = \sum_{m=0}^\infty \frac{(-1)^m \Gamma(\delta+1)}{m! \Gamma(\delta-m+1)} \{1 + \kappa x\}^{-\lambda m}$$

Hence

$$\begin{aligned} E(x^r) &= \frac{\vartheta \eta \kappa \lambda}{(1 - \{1 + \eta\}^{-\vartheta})} \sum_{\delta=0}^\infty \frac{\Gamma(\vartheta+1+\delta)}{\delta! \Gamma(\vartheta+1)} (-\eta)^\delta \sum_{m=0}^\infty \frac{(-1)^m \Gamma(\delta+1)}{m! \Gamma(\delta-m+1)} \int_0^\infty x^r (1 + \kappa x)^{-(\lambda+1)} (1 + \kappa x)^{-\lambda m} dx \\ &= \frac{\vartheta \eta \kappa \lambda}{(1 - \{1 + \eta\}^{-\vartheta})} \sum_{\delta=0}^\infty \frac{\Gamma(\vartheta+1+\delta)}{\delta! \Gamma(\vartheta+1)} (-\eta)^\delta \sum_{m=0}^\infty \frac{(-1)^m \Gamma(\delta+1)}{m! \Gamma(\delta-m+1)} \int_0^\infty x^r \{1 + \kappa x\}^{-(m+1)\lambda-1} dx \end{aligned}$$

$$\text{Let } y = 1 + \kappa x \rightarrow y - 1 = \kappa x \rightarrow \frac{y-1}{\kappa} = x \rightarrow dx = \frac{1}{\kappa} dy,$$

Thus:

$$\begin{aligned} E(x^r) &= \frac{\vartheta \eta \kappa \lambda}{[1 - \{1 + \eta\}^{-\vartheta}]} \sum_{\delta=0}^\infty \frac{\Gamma(\vartheta+1+\delta)}{\delta! \Gamma(\vartheta+1)} [-\eta]^\delta \sum_{m=0}^\infty \frac{(-1)^m \Gamma(\delta+1)}{m! \Gamma(\delta-m+1)} \frac{(-1)^r}{\kappa^{r+1}} \\ &\quad \int_1^\infty y^{-(m+1)\lambda-1} \{1 - y\}^r dy \end{aligned}$$

$$E(x^r) = \frac{\vartheta \eta \lambda (-1)^r}{\kappa^r [1 - \{1 + \eta\}^{-\vartheta}]} \sum_{\delta=0}^\infty \frac{\Gamma(\vartheta+1+\delta)}{\delta! \Gamma(\vartheta+1)} [-\eta]^\delta \sum_{m=0}^\infty \frac{(-1)^m \Gamma(\delta+1)}{m! \Gamma(\delta-m+1)}$$

$$\left(\int_0^\infty y^{-(m+1)\lambda-1} \{1 - y\}^r dy - \int_0^1 y^{-(m+1)\lambda-1} \{1 - y\}^r dy \right) \quad (2.6)$$

$$\text{Let } I_1 = \int_0^\infty y^{-(m+1)\lambda-1} (1 - y)^r dy ,$$

$$I_2 = \int_0^1 y^{-(m+1)\lambda-1} (1 - y)^r dy$$

Now , finding the integrations of I_1 and I_2 , respectively , and substituting the results in equation (2.6) , we get

$$I_1 = \frac{1}{2} \int_{-\infty}^\infty y^{-(m+1)\lambda-1} (1 - y)^r dy$$

Now ,simplifying $(1 - y)^r$ by using Eq (2.5)

we get :

$$I_1 = \frac{1}{2} \int_{-\infty}^\infty y^{-(m+1)\lambda-1} \sum_{j=0}^\infty (-1)^j \frac{\Gamma(r+1)}{j! \Gamma(r-j+1)} y^j dy$$

$$I_1 = \frac{1}{2} \sum_{j=0}^\infty (-1)^j \frac{\Gamma(r+1)}{j! \Gamma(r-j+1)} \int_{-\infty}^\infty y^{-(m+1)\lambda-1+j} dy = 0$$

$$I_2 = \int_0^1 y^{-(m+1)\lambda-1} (1 - y)^{(r+1)-1} dy$$

$$I_2 = B(-(m+1)\lambda, r+1)$$

$$I_2 = \frac{\Gamma(-(m+1)\lambda) \Gamma(r+1)}{\Gamma(-(m+1)\lambda + r + 1)}$$

$$\begin{aligned} E(x^r) &= \frac{\vartheta \eta \lambda (-1)^{r+1}}{\kappa^r (1 - \{1 + \eta\}^{-\vartheta})} \sum_{\delta=0}^\infty \frac{\Gamma(\vartheta+1+\delta)}{\delta! \Gamma(\vartheta+1)} (-\eta)^\delta \\ &\quad \sum_{m=0}^\infty \frac{(-1)^m \Gamma(\delta+1)}{m! \Gamma(\delta-m+1)} \frac{\Gamma(-(m+1)\lambda) \Gamma(r+1)}{\Gamma(-(m+1)\lambda + r + 1)} \quad (2.7) \end{aligned}$$

The Characteristic function is given by:

$$\Psi_x(t) = E(e^{ixt})$$

$$\Psi_x(t) = E \left[\sum_{r=0}^\infty \frac{(it)^r}{r!} x^r \right] = \sum_{r=0}^\infty \frac{(it)^r}{r!} E(x^r)$$

$$\begin{aligned} \Psi_x(t) &= \frac{\vartheta \eta \lambda}{(1 - \{1 + \eta\}^{-\vartheta})} \sum_{r=0}^\infty \left(\frac{-it}{\kappa} \right)^r \left(\frac{-1}{r!} \right) \sum_{\delta=0}^\infty \frac{\Gamma(\vartheta+1+\delta) (-\eta)^\delta}{\delta! \Gamma(\vartheta+1)} \\ &\quad \sum_{m=0}^\infty \frac{(-1)^m \Gamma(\delta+1)}{m! \Gamma(\delta-m+1)} \frac{\Gamma(-(m+1)\lambda) \Gamma(r+1)}{\Gamma(-(m+1)\lambda + r + 1)} \end{aligned}$$

Then :

The mode and the median can be calculated respectively,

$$\frac{df(x)}{dx} = \frac{-\eta\vartheta\kappa\lambda}{(1-(1+\eta)^{-\vartheta})} \cdot (\vartheta+1)(1+\kappa x)^{-2(\lambda+1)}\eta\lambda\kappa(1+\eta\{1-(1+\kappa x)^{-\lambda}\})^{-(\vartheta+2)} + \frac{\eta\vartheta\kappa\lambda}{(1-(1+\eta)^{-\vartheta})}\kappa(-\lambda-1)(1+\kappa x)^{-(\lambda+2)}(1+\eta\{1-(1+\kappa x)^{-\lambda}\})^{-(\vartheta+1)} = 0$$

$$\begin{aligned} & \frac{-\eta\vartheta\kappa\lambda}{(1-(1+\eta)^{-\vartheta})} \cdot (\vartheta+1)(1+\kappa x)^{-2(\lambda+1)}\eta\lambda\kappa(1+\eta\{1-(1+\kappa x)^{-\lambda}\})^{-(\vartheta+2)} \\ &= \frac{\eta\vartheta\kappa\lambda}{(1-(1+\eta)^{-\vartheta})}\kappa(-\lambda-1)(1+\kappa x)^{-(\lambda+2)}(1+\eta\{1-(1+\kappa x)^{-\lambda}\})^{-(\vartheta+1)} \\ & (\vartheta+1)(1+\kappa x)^{-\lambda} \cdot \eta\lambda\kappa(1+\eta\{1-(1+\kappa x)^{-\lambda}\})^{-1} \\ &= -\kappa\lambda - \kappa \end{aligned}$$

$$\frac{(1+\kappa x)^{-\lambda}}{(1+\eta\{1-(1+\kappa x)^{-\lambda}\})} = \frac{(-\lambda-1)}{(\eta\lambda\vartheta + \eta\lambda)}$$

$$\eta\lambda\vartheta(1+\kappa x)^{-\lambda} + \eta\lambda(1+\kappa x)^{-\lambda} = -\lambda(1+\eta\{1-(1+\kappa x)^{-\lambda}\}) - (1+\eta\{1-(1+\kappa x)^{-\lambda}\})$$

$$\eta\lambda\vartheta(1+\kappa x)^{-\lambda} - \eta(1+\kappa x)^{-\lambda} = -\lambda - \lambda\eta - 1 - \eta$$

$$(1+\kappa x)^{-\lambda} = \frac{-\lambda - \lambda\eta - 1 - \eta}{(\eta\lambda\vartheta - \eta)} = \frac{-(\lambda + \lambda\eta + 1 + \eta)}{\eta(\lambda\vartheta - 1)} = \frac{-\{\lambda(1+\eta) + (1+\eta)\}}{\eta(\lambda\vartheta - 1)}$$

$$(1+\kappa x)^{-\lambda} = \frac{-[(1+\eta)(\lambda+1)]}{\eta(\lambda\vartheta - 1)}$$

$$x_{Mode} = \frac{\left[\left(\left(\frac{-[(1+\eta)(\lambda+1)]}{\eta(\lambda\vartheta - 1)} \right) \right)^{\frac{-1}{\lambda}} - 1 \right]}{\kappa} \quad (2.12)$$

$$F(x) = \frac{1}{2}$$

$$\frac{(1 - [1 + \eta\{1 - (1 + \kappa x)^{-\lambda}\}]^{-\vartheta})}{(1 - (1 + \eta)^{-\vartheta})} = \frac{1}{2}$$

Then

$$1 - \frac{\{1 - (1 + \eta)^{-\vartheta}\}}{2} = (1 + \eta\{1 - (1 + \kappa x)^{-\lambda}\})^{-\vartheta}$$

$$\left[\frac{1 + (1 + \eta)^{-\vartheta}}{2} \right]^{\frac{-1}{\vartheta}} - 1 = \eta\{1 - (1 + \kappa x)^{-\lambda}\}$$

$$\frac{\left\{ \left[\frac{1 + (1 + \eta)^{-\vartheta}}{2} \right]^{\frac{-1}{\vartheta}} - 1 \right\}}{\eta} = 1 - (1 + \kappa x)^{-\lambda}$$

$$1 + \kappa x = \left(1 - \frac{\left\{ \left[\frac{1 + (1 + \eta)^{-\vartheta}}{2} \right]^{\frac{-1}{\vartheta}} - 1 \right\}}{\eta} \right)^{\frac{-1}{\lambda}}$$

$$x_{median} = \frac{\left[\left(1 - \frac{\left\{ \left(\frac{1 + (1 + \eta)^{-\vartheta}}{2} \right)^{\frac{-1}{\vartheta}} - 1 \right\}}{\eta} \right)^{\frac{-1}{\lambda}} \right]}{\kappa} \quad (2.13)$$

The quantile function x_q can be find it as follows ,

$$q = P(X \leq x_q) = F_x(x_q) = \frac{[1 - (1 + \eta\{1 - (1 + \kappa x)^{-\lambda}\})^{-\vartheta}]}{[1 - (1 + \eta)^{-\vartheta}]}$$

$$0 < q < 1, x_q > 0$$

$$1 - (1 + \eta\{1 - (1 + \kappa x)^{-\lambda}\})^{-\vartheta} = F_x(x_q) \cdot [1 - (1 + \eta)^{-\vartheta}]$$

$$(1 + \eta\{1 - (1 + \kappa x)^{-\lambda}\})^{-\vartheta} = 1 - F_x(x_q)[1 - (1 + \eta)^{-\vartheta}]$$

$$(1 - F_x(x_q)[1 - (1 + \eta)^{-\vartheta}])^{\frac{-1}{\vartheta}} - 1 = \eta\{1 - (1 + \kappa x)^{-\lambda}\}$$

$$x_q = F^{-1}(q) = \frac{\left[\left\{ 1 - \left(\frac{(1 - F_x(x_q)(1 - (1 + \eta)^{-\vartheta}))^{\frac{-1}{\vartheta}} - 1}{\eta} \right) \right\}^{\frac{-1}{\lambda}} \right]}{\kappa} \quad (2.14)$$

By the inverse transform technique, we can create (r.v.) for ([0,1] TLLD) as follows,

$$x_p = \frac{\left[\left\{ 1 - \left(\frac{(1 - U \cdot (1 - (1 + \eta)^{-\vartheta}))^{\frac{-1}{\vartheta}} - 1}{\eta} \right) \right\}^{\frac{-1}{\lambda}} \right]}{\kappa}, 0 \leq U \leq 1$$

III. STRESS STRENGTH MODEL OF THE [0,1]TLLD DISTRIBUTION.

The stress strength is obtained through the formula as follows

$$R = P(y < x) = \int_0^\infty f_x(x) F_Y(x) dx$$

Now , Let $X \sim [0,1]$ TLLD $(\eta, \vartheta, \kappa, \lambda)$ and

$Y \sim [0,1]$ TLLD $(\eta_1, \vartheta_1, \kappa_1, \lambda_1)$

respectively ,

$$R = P(y < x) = \int_0^\infty f_x(x) F_Y(x) dx$$

$$R = \int_0^\infty \frac{\vartheta \eta \kappa \lambda}{(1-(1+\eta)^{-\vartheta})} (1+\kappa x)^{-(\lambda+1)} (1+\eta \{1 - (1+\kappa x)^{-\lambda}\})^{-(\vartheta+1)} \frac{[1-(1+\eta_1 \{1-(1+\kappa_1 x)^{-\lambda_1}\})^{-\vartheta_1}]}{(1-(1+\eta_1)^{-\vartheta_1})} dx$$

Simplifying :

$(1+\eta \{1 - (1+\kappa x)^{-\lambda}\})^{-(\vartheta+1)}$ by using the series expansion ,

$$(1-z)^{-k} = \sum_{s=0}^\infty \frac{\Gamma(k+s)}{s! \Gamma(k)} z^s \quad |z| < 1, k > 0, \text{ we obtain}$$

$$(1+\eta \{1 - (1+\kappa x)^{-\lambda}\})^{-(\vartheta+1)} = (1+\eta - \eta(1+\kappa x)^{-\lambda})^{-(\vartheta+1)}$$

$$(1 - \{\eta(1+\kappa x)^{-\lambda} - \eta\})^{-(\vartheta+1)} = \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} (\eta(1+\kappa x)^{-\lambda} - \eta)^s$$

$$(1 - \{\eta(1+\kappa x)^{-\lambda} - \eta\})^{-(\vartheta+1)} = \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} (-\eta)^s (1 - (1+\kappa x)^{-\lambda})^s$$

Then:

$$SR = \frac{\vartheta \eta \kappa \lambda}{(1-(1+\eta)^{-\vartheta})} \frac{1}{(1-(1+\eta_1)^{-\vartheta_1})} \int_0^\infty (1 + \kappa x)^{-(\lambda+1)} \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} (-\eta)^s (1 - (1+\kappa x)^{-\lambda})^s \{1 - (1+\eta_1 \{1 - (1+\kappa_1 x)^{-\lambda_1}\})^{-\vartheta_1}\} dx$$

$$R = \frac{\vartheta \eta \kappa \lambda}{[1-(1+\eta)^{-\vartheta}]} \frac{1}{[1-(1+\eta_1)^{-\vartheta_1}]} \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} [-\eta]^s \int_0^\infty \{1 + \kappa x\}^{-\lambda} [1 - \{1 + \kappa x\}^{-\lambda}]^s (1 - \{1 + \eta_1 [1 - \{1 + \kappa_1 x\}^{-\lambda_1}\}]^{-\vartheta_1}) dx$$

Also, simplifying :

$$(1 - (1 + \kappa x)^{-\lambda})^s \text{ by using}$$

$$(1-z)^b = \sum_{u=0}^\infty (-1)^u \frac{\Gamma(b+1)}{u! \Gamma(b-u+1)} z^u ,$$

$|z| < 1, b > 0$, we get

$$[1 - \{1 + \kappa x\}^{-\lambda}]^s = \sum_{u=0}^\infty \frac{(-1)^u \Gamma(s+1)}{u! \Gamma(s-u+1)} ((1 + \kappa x))^{-\lambda u}$$

Therefore,

$$= \frac{\vartheta \eta \kappa \lambda}{(1-(1+\eta)^{-\vartheta})} \frac{1}{(1-(1+\eta_1)^{-\vartheta_1})} \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} (-\eta)^s \sum_{u=0}^\infty (-1)^u \frac{\Gamma(s+1)}{u! \Gamma(s-u+1)}$$

$$\int_0^\infty (1 + \kappa x)^{-(u+1)\lambda} [1 - (1 + \eta_1 \{1 - (1 + \kappa_1 x)^{-\lambda_1}\})^{-\vartheta_1}]^{-\vartheta_1} dx$$

Using the same simplification approach:

$$(1 + \eta_1 \{1 - (1 + \kappa_1 x)^{-\lambda_1}\})^{-\vartheta_1} \text{ and using}$$

$$(1-z)^{-s} = \sum_{i=0}^\infty \frac{\Gamma(s+i)}{i! \Gamma(s)} z^i \quad |z| < 1, s > 0, \text{ gives}$$

$$(1 + \eta_1 \{1 - (1 + \kappa_1 x)^{-\lambda_1}\})^{-\vartheta_1} = \sum_{i=0}^\infty \frac{\Gamma(\vartheta_1+i)}{i! \Gamma(\vartheta_1)} \{\eta_1 (1 + \kappa_1 x)^{\lambda_1} - \eta_1\}^i$$

Then:

$$R = \frac{\vartheta \eta \kappa \lambda}{(1-(1+\eta)^{-\vartheta})} \frac{1}{(1-(1+\eta_1)^{-\vartheta_1})} \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} (-\eta)^s \sum_{u=0}^\infty (-1)^u \frac{\Gamma(s+1)}{u! \Gamma(s-u+1)}$$

$$\int_0^\infty (1 + \kappa x)^{-(u+1)\lambda} \left(1 - \sum_{i=0}^\infty \frac{\Gamma(\vartheta_1+i)}{i! \Gamma(\vartheta_1)} \{\eta_1 (1 + \kappa_1 x)^{\lambda_1} - \eta_1\}^i \right) dx$$

$$= \frac{\vartheta \eta \kappa \lambda}{(1-(1+\eta)^{-\vartheta})} \frac{1}{(1-(1+\eta_1)^{-\vartheta_1})} \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} (-\eta)^s \sum_{u=0}^\infty (-1)^u \frac{\Gamma(s+1)}{u! \Gamma(s-u+1)}$$

$$\int_0^\infty (1 + \kappa x)^{-(u+1)\lambda} dx - \frac{\vartheta \eta \kappa \lambda}{(1-(1+\eta)^{-\vartheta})} \frac{1}{(1-(1+\eta_1)^{-\vartheta_1})} \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} (-\eta)^s \sum_{u=0}^\infty (-1)^u \frac{\Gamma(s+1)}{u! \Gamma(s-u+1)}$$

$$\int_0^\infty (1 + \kappa x)^{-(u+1)\lambda} \left(\sum_{i=0}^\infty \frac{\Gamma(\vartheta_1+i)}{i! \Gamma(\vartheta_1)} \{\eta_1 (1 + \kappa_1 x)^{\lambda_1} - \eta_1\}^i \right) dx$$

$$= \frac{\vartheta \eta \kappa \lambda}{(1-(1+\eta)^{-\vartheta})} \frac{1}{(1-(1+\eta_1)^{-\vartheta_1})} \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} (-\eta)^s \sum_{u=0}^\infty (-1)^u \frac{\Gamma(s+1)}{u! \Gamma(s-u+1)}$$

$$\int_0^\infty (1 + \kappa x)^{-(u+1)\lambda} dx - \frac{\vartheta \eta \kappa \lambda}{(1-(1+\eta)^{-\vartheta})} \frac{1}{(1-(1+\eta_1)^{-\vartheta_1})} \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} (-\eta)^s \sum_{u=0}^\infty (-1)^u \frac{\Gamma(s+1)}{u! \Gamma(s-u+1)}$$

$$\sum_{i=0}^\infty \frac{\Gamma(\vartheta_1+i)}{i! \Gamma(\vartheta_1)} \int_0^\infty (1 + \kappa x)^{-(u+1)\lambda} (-\eta_1)^i \{1 - (1 + \kappa_1 x)^{-\vartheta_1}\}^i dx$$

Simplifying $\{1 - (1 + \kappa_1 x)^{-\vartheta_1}\}^i$ by using $(1-z)^m =$

$$\sum_{n=0}^\infty (-1)^n \frac{\Gamma(m+1)}{n! \Gamma(m-n+1)} z^n ,$$

$|z| < 1, b > 0$, gives

$$\{1 - (1 + \kappa_1 x)^{-\vartheta_1}\}^i = \sum_{n=0}^\infty (-1)^n \frac{\Gamma(i+1)}{i! \Gamma(i-n+1)} (1 + \kappa_1 x)^{-\lambda_1 n}$$

$$= \frac{\vartheta \eta \kappa \lambda}{(1-(1+\eta)^{-\vartheta})} \frac{1}{(1-(1+\eta_1)^{-\vartheta_1})} \sum_{s=0}^\infty \frac{\Gamma(\vartheta+1+s)}{s! \Gamma(\vartheta+1)} (-\eta)^s$$

$$\sum_{u=0}^\infty (-1)^u \frac{\Gamma(s+1)}{u! \Gamma(s-u+1)} \int_0^\infty (1 + \kappa x)^{-(u+1)\lambda} dx$$

$$-\frac{\vartheta\eta\kappa\lambda}{(1-(1+\eta)^{-\vartheta})}\frac{1}{(1-(1+\eta_1)^{-\vartheta_1})}\sum_{s=0}^{\infty}\frac{\Gamma(\vartheta+1+s)}{s!\Gamma(\vartheta+1)}(-\eta)^s$$

$$\sum_{u=0}^{\infty}(-1)^u\frac{\Gamma(s+1)}{u!\Gamma(s-u+1)}\sum_{i=0}^{\infty}\frac{\Gamma(\vartheta_1+i)}{i!\Gamma(\vartheta_1)}$$

$$\int_0^{\infty}(1+\kappa x)^{-(u+1)\lambda}(-\eta_1)^i\sum_{n=0}^{\infty}(-1)^n\frac{\Gamma(i+1)}{n!\Gamma(i-n+1)}(1+\kappa_1 x)^{-\lambda_1 i}dx$$

Since $(1+\kappa_1 x)^{-\lambda_1 i} = (1 - (-\kappa_1 x))^{-\lambda_1 i}$ by using

$$(1-z)^{-k} = \sum_{v=0}^{\infty} \frac{\Gamma(k+v)}{v!\Gamma(k)} z^v \quad |z| < 1, k > 0,$$

gives,

$$(1 - (-\kappa_1 x))^{-\lambda_1 i} = \sum_{v=0}^{\infty} \frac{\Gamma(\lambda_1 i + v)}{v!\Gamma(\lambda_1 i)} (-\kappa_1 x)^v$$

$$= \sum_{v=0}^{\infty} \frac{\Gamma(\lambda_1 i + v)}{v!\Gamma(\lambda_1 i)} (-\kappa_1)^v x^v$$

Then:

$$R =$$

$$\frac{\vartheta\eta\kappa\lambda}{(1-(1+\eta)^{-\vartheta})}\frac{1}{(1-(1+\eta_1)^{-\vartheta_1})}\sum_{s=0}^{\infty}\frac{\Gamma(\vartheta+1+s)}{s!\Gamma(\vartheta+1)}(-\eta)^s\sum_{u=0}^{\infty}(-1)^u\frac{\Gamma(s+1)}{u!\Gamma(s-u+1)}$$

$$\int_0^{\infty}(1+\kappa x)^{-(u+1)\lambda}dx -$$

$$\frac{\vartheta\eta\kappa\lambda}{(1-(1+\eta)^{-\vartheta})}\frac{1}{(1-(1+\eta_1)^{-\vartheta_1})}\sum_{s=0}^{\infty}\frac{\Gamma(\vartheta+1+s)}{s!\Gamma(\vartheta+1)}(-\eta)^s\sum_{u=0}^{\infty}(-1)^u\frac{\Gamma(s+1)}{u!\Gamma(s-u+1)}$$

$$\sum_{i=0}^{\infty}\frac{\Gamma(\vartheta_1+i)}{i!\Gamma(\vartheta_1)}(-\eta_1)^i\sum_{n=0}^{\infty}(-1)^n\frac{\Gamma(i+1)}{n!\Gamma(i-n+1)}\sum_{v=0}^{\infty}\frac{\Gamma(\lambda_1 i+v)}{v!\Gamma(\lambda_1 i)}(-\kappa_1)^v\int_0^{\infty}x^v(1+\kappa x)^{-(u+1)\lambda}dx$$

$$(3.1)$$

Let $I_1 = \int_0^{\infty}(1+\kappa x)^{-(u+1)\lambda}dx$ and

$$I_2 = \int_0^{\infty}x^v(1+\kappa x)^{-(u+1)\lambda}dx$$

Now, finding the integrations of I_1 and I_2 , respectively, then replacing the results in (3.1), gives

Let $y = 1 + \kappa x \rightarrow y - 1 = \kappa x \rightarrow \frac{y-1}{\kappa} = x \rightarrow \frac{1}{\kappa} dy = dx$, therefore,

$$I_1 = \int_1^{\infty} y^{-(u+1)\lambda} \frac{1}{\kappa} dy = \frac{1}{\kappa} \int_1^{\infty} y^{-(u+1)\lambda} dy = \frac{-1}{\kappa}$$

$$I_2 = \int_1^{\infty} \left(\frac{y-1}{\kappa}\right)^v y^{-(u+1)\lambda} \frac{1}{\kappa} dy = \frac{1}{\kappa^{v+1}} \int_1^{\infty} y^{-(u+1)\lambda} (y-1)^v dy$$

$$I_2 = \frac{1}{\kappa^{v+1}} \left\{ \int_0^{\infty} y^{-(u+1)\lambda} (y-1)^v dy - \int_0^1 y^{-(u+1)\lambda} (y-1)^v dy \right\}$$

(3.2)

Now, Let

$$A_1 = \int_0^{\infty} y^{-(u+1)\lambda} (y-1)^v dy \text{ and}$$

$$A_2 = \int_0^1 y^{-(u+1)\lambda} (y-1)^v dy$$

Again, finding the integrations of A_1 and A_2 , respectively, then replacing the results in (3.2) gives

$$A_1 = \int_0^{\infty} y^{-(u+1)\lambda} (y-1)^v dy$$

$$= \frac{1}{2} \int_{-\infty}^{\infty} y^{-(u+1)\lambda} (y-1)^v dy$$

Using the same way, simplifying $(1-y)^v$ by using Eq.(2.5)

we obtain

$$A_1 = \frac{1}{2} \int_{-\infty}^{\infty} y^{-(u+1)\lambda} \sum_{m=0}^{\infty} (-1)^m \frac{\Gamma(v+1)}{m!\Gamma(v-m+1)} y^m dy$$

$$A_1 = \frac{1}{2} \sum_{m=0}^{\infty} (-1)^m \frac{\Gamma(v+1)}{m!\Gamma(v-m+1)} \int_{-\infty}^{\infty} y^{-(u+1)\lambda+m} dy$$

$$= 0$$

$$A_2 = \int_0^1 y^{-(u+1)\lambda+1-1} (1-y)^{(v+1)-1} dy$$

$$A_2 = B(-u\lambda - \lambda + 1, v+1) = \frac{\Gamma(-u\lambda - \lambda + 1) \Gamma(v+1)}{\Gamma(-u\lambda - \lambda + v+2)}$$

$$\text{Therefore } I_2 = \frac{-1}{\kappa^{v+1}} \frac{\Gamma(-u\lambda - \lambda + 1) \Gamma(v+1)}{\Gamma(-u\lambda - \lambda + v+2)}$$

Hence,

$$R =$$

$$\frac{-\vartheta\eta\lambda}{(1-(1+\eta)^{-\vartheta})}\frac{1}{(1-(1+\eta_1)^{-\vartheta_1})}\sum_{s=0}^{\infty}\frac{\Gamma(\vartheta+1+s)}{s!\Gamma(\vartheta+1)}(-\eta)^s\sum_{u=0}^{\infty}(-1)^u\frac{\Gamma(s+1)}{u!\Gamma(s-u+1)}$$

$$+\frac{\vartheta\eta\lambda}{(1-(1+\eta)^{-\vartheta})}\frac{1}{(1-(1+\eta_1)^{-\vartheta_1})}\sum_{s=0}^{\infty}\frac{\Gamma(\vartheta+1+s)}{s!\Gamma(\vartheta+1)}(-\eta)^s\sum_{u=0}^{\infty}(-1)^u\frac{\Gamma(s+1)}{u!\Gamma(s-u+1)}$$

$$\sum_{i=0}^{\infty}\frac{\Gamma(\vartheta_1+i)}{i!\Gamma(\vartheta_1)}(-\eta_1)^i\sum_{n=0}^{\infty}(-1)^n\frac{\Gamma(i+1)}{n!\Gamma(i-n+1)}\sum_{v=0}^{\infty}\frac{\Gamma(\lambda_1 i+v)}{v!\Gamma(\lambda_1 i)}$$

$$\frac{(-\kappa_1)^v \Gamma(-u\lambda - \lambda + 1) \Gamma(v+1)}{\kappa^v \Gamma(-u\lambda - \lambda + v+2)} \quad (3.3)$$

IV. CONCLUSIONS

In reality , we suggested ([0,1] TLLD) distribution dependent on ([0,1] TLD) distribution and It's mathematical properties have been obtained and also , stress-strength reliability.

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